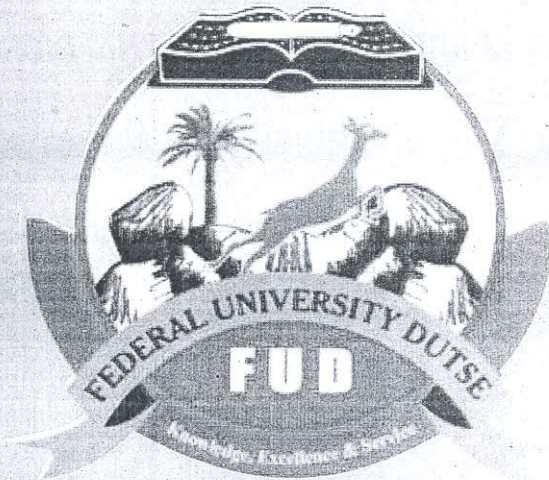


26



Dutse Journal of Humanities and Social Sciences

Volume 1,
September, 2014



Dutse Journal of Humanities and Social Sciences

Vol. 1, September, 2014

Editor-in-Chief Salisu A. Abdullahi
Editor Amina A. Bashir

No part of this publication may be reproduced, stored in a retrieval system or transmitted and distributed in any form or by any means without the prior permission of the copyright owner.

© Federal University Dutse, 2014
All Rights Reserved
djhss@fud.edu.ng
url:library.fud.edu.ng

Printed by
Ahmadu Bello University Press Limited,
Zaria, Kaduna State, Nigeria.
Tel.: 08065949711.
E-mail: abupresslimited2005@yahoo.co.uk;
abupress2013@gmail.com
Website: www.abupress.org

Editorial

Editorial Board:

Salisu A. Abdullahi (Sociology)	Chairman
Amina A. Bashir (English)	Secretary
Ibrahim A. Kiyawa (Economics)	Member
Habu M. Fagge (Political Science)	Member

Editorial Advisers:

B. Antia	Department of Languages and Linguistics University of Cape Town, South Africa
R. W. Morgan	Department of Sociology, Cleveland State University, USA
J. A. Falola	Department of Geography Bayero University Kano
O. Soyombo	Department of Sociology University of Lagos Nigeria
C. Bassey	Department of Political Science, University of Calabar, Nigeria
D. Danju	Department of Economics, Eastern Mediterranean University, Cyprus

Contents

The Phonological Alternations of Individual Graphemes in Received Pronunciation (RP) in English Amina A. Bashir and Ozim M. Chika...	1
A Pragmatic Analysis of Visuals in Select Advertisements of the New Media Charles G. Doke, Eneojoh J. Amodu & Monday A. Bello Eje...	19
A Pragmatic Analysis of Information Packaging in the Creative Language Content of Alcohol Advertising Eneojoh J. Amodu...	43
Development Planning in Nigeria: The Consequences of Ideological Shift Ibrahim A. Kiyawa...	55
The Determinants of Technical Efficiency Among Selected Small Scale Farmers Adama J. Joseph...	89 *
The Impact of HIV/AIDS on Capital Formation in Nigeria Angbas J. Avreson...	105
The Election Management and Legitimacy Crisis in Nigeria Hassan M. Babatunde...	125
Diversity and National Integration Mohammed B. Umar and Lamidi K. Olaniyi...	149
The Transformation of the Nigerian Military Adebola R. Bakare	165

Durkheim, Weber and Geertz on Religion: Implications for Theorizing in Africa Mustapha H. Kurfi...	191
The Socio-Economic Implications of Kidnapping John E. Gyong and Danjuma S. Asaka... ...	209
Drug Abuse by Youth in Kano Mohammed A. Haruna...	243
Some Remarks Regarding Prison Violence in Nigeria Salisu A. Abdullahi...	259
Sexual Violence Awareness Amongst University Students and Its Implication for Reproductive Health Care Bene Nnabugwu-Otesanya...	275
Sexuality and Sexual Behaviour Among Adolescent Girls and Younger Women Bola L. Solanke...	291
Mothers' Media Exposure and Children Immunization Coverage in Lagos John L. Oyefara	313

The Determinants of Technical Efficiency among Selected Small Scale Farmers in Kogi State

Adama I. Joseph

Department of Economics,
Ahmadu Bello University, Zaria

Abstract

The purpose of this study was to analyze the determinants of the technical efficiency of small scale farmers in Kogi State of Nigeria and the effect of policy variable changes on technical efficiency, using the stochastic frontier methodology. Results of the analysis revealed that the farmers have an average farm size of 1.72 hectares. It is also indicated that both family and hired labour were expansively used in farm production. The analysis shows a spacious variation in the estimated technical efficiencies within the range of 0.19 and 0.93 and a mean value of 0.64, indicating ample opportunity for enhancement in technical efficiency. The results of simulation of policy variables prove that the level of technical efficiency would extensively increase with rising levels of farming experiences and educational level among the rural farmers.

Keywords: technical efficiency, small scale farmers, stochastic frontier, production function.

Introduction

Kogi State agriculture is predominantly the work of small scale farmers who produce the bulk of food requirements in the State. Despite their distinctive and fundamental position, the small holder farmers belong to the poorest section of the population and, therefore, cannot spend much on their farms. The vicious circle of poverty among these farmers has led to the low performance of the agricultural sector. While substantial efforts have been made to raise the production and productivity of these farmers so as to achieve food security, such efforts have had negative implications for the environment. As the population

density increases, farmers must produce even more food than before. With the population increases of today, people are being pushed to new lands and many into marginal lands. One of the enormous challenges in the drive to increase food to feed the growing population will be to raise productivity and efficiency in the agricultural sector, especially because Nigeria's rapid population growth has outstripped the nation's capacity to grow food. Similarly, Ojo (1993) wrote that, from 1991 - 2000, the Kogi state population grew by 2.1% a year, while agricultural production lagged far behind, growing at just 2.5% a year. Also, given the various agricultural programmes and policies implemented in the state over the years to raise farmers' efficiency and productivity, it is imperative to quantitatively measure the current level and determinants of the technical efficiency and policy options available for raising the present level of efficiency. Coupled with that, efficiency in production is directly related to the overall productivity of the agricultural sector. From the foregoing, there is the crucial need to raise agricultural growth; as such, growth is the most efficient means of alleviating poverty and protecting the environment. Therefore, raising productivity per area of land is the key to effectively addressing the challenges of achieving food security in the state, as most cultivable land has already been brought under cultivation. In areas where wide expanses of cultivable land are still available, physical and technological constraints prevent large-scale conversion of potentially cultivable land.

The main objective of this study is to estimate technical efficiency and identify the factors that explain variations in technical efficiency. The study has two specific objectives. First, it examines the impact of technology adoption on the technical efficiency of smallholder farmers in Kogi State. Secondly, it determines the relative role of farmer education and the use of farm tools.

Conceptual Underpinning

In conceptualizing efficiency, Akpakpan (1999) defined efficiency as the objective of getting the most from an undertaking. A resource is said to be efficiently used if its

marginal variable cost (MVC) exceeds its marginal factor cost (MFC). Similarly, Farrel (1957) explained the term "technical efficiency" as the ability to produce the highest level of output given a bundle of resources. On the other hand, technical inefficiency depicts a situation in which actual and observed output from a given input mix is less than the maximum possible. The modern theory of efficiency dates back to the pioneering work of Farrel (1957), who drew extensively from the earlier works of Debreu (1951) and Koopmans (1951) to define a simple measure of firm efficiency, which could account for multiple inputs. Farrel identified two components of firm efficiency-technical and allocative and the combination of these two components provides a measure of economic efficiency. Thus, technical efficiency, the main issue in this study, can be measured either as input conserving oriented technical efficiency or output expanding oriented technical efficiency. Output expanding oriented technical efficiency is the ratio of observed to maximum feasible output, conditional on technical and observed input usage (Udoh, 2005).

Theoretical Review

There are a number of different theoretical approaches that attempt to explain the efficiency of farmers in crop production. As a component of productive efficiency, technical efficiency is derived from the production function. Productive efficiency consists of technical efficiency and allocative or factor price efficiency. Productive efficiency represents the efficient resource input mix for any given output that minimizes the cost of producing that level of output or equivalently the combination of inputs that for a given monetary outlay maximizes the level of production. Technical efficiency reflects the ability of a firm to maximize output for a given set of resource inputs, while allocative (factor price) efficiency reflects the ability of the firm to use the inputs in optimal proportions given their respective prices and the production technology. Developments in cost and production frontiers are attempts to measure productive efficiency, as proposed by Udoh (2005). The frontier defines the limit to a range of possible observed production (cost) levels and

identifies the extent to which the firm lies below (above) the frontier.

Estimating Technical Efficiency

The literature suggests several alternative approaches to measuring productive efficiency grouped into non-parametric frontiers and parametric frontiers. Nonparametric frontiers do not impose a functional form on the production frontiers and do not make assumptions about the error term. These have used linear programming approaches; the most popular non-parametric approach has been the data envelopment analysis. Parametric frontier approaches impose a functional form on the production function and make assumptions about the data. The most common functional forms include the Cobb–Douglas, constant elasticity of substitution and translog production functions. The other distinction is between deterministic and stochastic frontiers. Deterministic frontiers assume that all the deviations from the frontier are a result of firms' inefficiency, while stochastic frontiers assume that part of the deviation from the frontier is due to random events (reflecting measurement errors and statistical noise) and part is due to firm specific inefficiency. The stochastic frontier approach, unlike the other parametric frontier measures, makes allowance for stochastic errors arising from statistical noise or measurement errors. The stochastic frontier model decomposes the error term into a two-sided random error that captures the random effects outside the control of the firm (the decision making unit) and the one-sided efficiency component. The model was first proposed by Aigner et al(1977).

The literature suggests two methodological approaches for analysing the sources of technical efficiency based on stochastic production functions. The first approach is the two-stage estimation procedure in which, first, the stochastic production function is estimated, from which efficiency scores are derived. In the second stage, the derived efficiency scores are regressed on explanatory variables, using ordinary least square methods or tobit regression. This approach has been criticized on grounds that the firm's knowledge of its level of technical

inefficiency affects its input choices; hence, inefficiency may be dependent on the explanatory variables. The second approach advocates a simultaneous estimation approach, as in Battese and Coelli (1995), in which the inefficiency effects are expressed as an explicit function of a vector of farm-specific variables.

The technical inefficiency effects are expressed as $u_j = Z_j \delta$ Where, for farm j , z is a vector of observable explanatory variables and δ is a vector of unknown parameters. Thus, the parameters of the frontier production function are simultaneously estimated with those of an inefficiency model, in which the technical inefficiency effects are specified as a function of other variables. The one-stage simultaneous approach is also implemented in FRONTIER and in addition to the basic parameters the programme also provides coefficients for the technical inefficiency model. Several factors, including socioeconomic and demographic factors, plot-level characteristics, environmental factors, and non-physical factors, are likely to affect the efficiency of smallholder farmers. Moreover, Parikh et al. (1995), using stochastic cost frontiers in Pakistani agriculture in a two-stage estimation procedure, found that education, the number of working animals, credit per acre and the number of extension visits significantly increase cost efficiency, while large land holding size and subsistence significantly decrease it.

Existing Empirical Studies in Africa

The literature on productive or technical efficiency in African agriculture is emerging. Globally, however, there is a wide body of empirical research on the economic efficiency of farmers in both developed and developing countries (See Battese and Coelli, 1995). While the empirical literature on the efficiency of farmers is vast in developed countries and Asian economies, few studies focus on African agriculture. For instance, Heshmati and Mulugeta (1996) estimated the technical efficiency of Ugandan *matoke* producing farms and found that they face decreasing returns to scale with a mean technical efficiency of 65%. On the other hand, they found no significant variation in technical efficiency with respect to farm sizes. Nor

did they identify the various sources of technical efficiency among *matoke*-producing farmers. Also, Seyoum et al. (1998) consider the technical efficiency and productivity of maize producers in Ethiopia and compare the performance of farmers within and outside the programme of technology demonstration. Using Cobb–Douglas stochastic production functions, their empirical results show that farmers who participate in the programme are more technically efficient with a mean technical efficiency equal to 94% compared with those outside the project whose mean efficiency equalled 79%.

Similarly, in Ethiopia Weir (1999) investigated the effects of education on farmer productivity of cereal crops using average and stochastic production functions. This study found substantial internal benefits of schooling for farmer productivity in terms of efficiency gains, but found a threshold effect that implied that at least four years of schooling are required to lead to significant effects on farm level technical efficiency. Using different specifications, average technical efficiencies range between 0.44 and 0.56, and raising education from zero to four years in the household leads to a 15% increase in technical efficiency. Moreover, the study discovered that average schooling in the villages (the external benefits of schooling) improves technical efficiency. The impact of education externalities on production and the technical efficiency of farmers in rural Ethiopia is the subject of Weir and Knight (2000). They found evidence that the source of externalities to schooling is in the adoption and spread of innovations that shift out the production frontier. Mean technical efficiencies of cereal crop farmers are 0.55. A unit increase in years of schooling increases technical efficiency by 2.1 percentage points. One limitation of the Weir (1999) and Weir and Knight (2000) studies is that they investigate the levels of schooling as the only source of technical efficiency. Using data envelopment analysis, Townsend et al. (1998) investigate the relationships among farm size, returns to scale and productivity for wine producers in South Africa. They found that most farmers operate under constant returns to scale, but the inverse relationship between farm size and productivity is weak. Mochebelele and Winter-

Nelson (2000) assess the impact of labour migration on the technical efficiency performance of farms in the rural economy of Lesotho. Using the stochastic production function (translog and Cobb–Douglas), the study found that households that send migrant labour to South African mines are more efficient than those that do not, with mean inefficiencies of 0.36 and 0.24, respectively. In addition, there is no statistical evidence that the size of the farm or the gender of the household head affects the efficiency of farmers. These authors conclude that remittances facilitate agricultural production, rather than substitute for it. Their study did not, however, consider the many other household characteristics that may affect technical efficiency, such as education, farmers' experience, access to credit facilities (capital) and advisory services, and the extent to which households that export labour receive remittances. The authors' interpretation that it is remittances that explain differences in technical efficiencies is based on the presumption that migrant labourers remit to their exporting households, and not on some measures of the extent of remittances. Sherlund et al. (2002) investigate the efficiency of smallholder rice farmers in Côte d'Ivoire while controlling for environmental factors that affect the production process. Apart from identifying factors that influence technical efficiencies, the study finds that the inclusion of environmental variables in the production function significantly changes the results: the estimated mean technical efficiencies increase from 0.76% to 36%. Binam et al. (2004) examine the factors influencing the technical efficiency of groundnut and maize farmers in Cameroon. They use a Cobb–Douglas production function to find mean technical efficiencies to be in the region of 73% and 77%. They also conclude that access to credit, social capital, distance from the road and extension services are important factors explaining the variations in technical efficiencies.

The Study Area

Kogi is one of the central states of Nigeria created on the 27th of August, 1991. The state is located between latitudes 60:33' and 80:44' N and longitudes 50:22' and 70: 49' E. It has a

total population of about 3.278 million with an average of about 228,964 farm families (Kogi state Min. of Agric, 1999). About 70% of the people live in the rural areas and mostly engage in one form of agricultural practice or the other. The average farm family in the state is made up of 7 people with an average farm size of about 2 hectares per farmer. Administratively, the state has twenty-one Local Government Areas (LGAs). Predominant agricultural crops cultivated in the state includes cassava, maize, cocoyam and yam. Most of the agricultural products are obtained from rain-fed irrigated crops farmed with a mixed cropping system.

Data for the Study

The selection of respondent farmers for this study was multistage. The first stage involves the selection of two LGAs from the three senatorial districts. The local governments in the state were divided into villages and one village was selected from each of the local governments. The second stage of the sampling procedure demands the random selection of villages proportionate to the village population of the identified LGAs. From each selected village, 40 smallholder farmers were interviewed, making a total of 240 sample farmers in all. Major Production resources used by the farmers were grouped into five. These are: land, labour, seed, implements and chemicals. Majority of the farmers relied mostly on these productive resources. The study measured land in hectares; and human labour in man days (family and hired labour). Seed chemical and implements were each measured in quantity and the price of the resources used.

Model Specification

This study uses the stochastic frontier production function. The model has the advantage in that it allows the simultaneous estimation of individual technical efficiency of the respondent farmers, as well as the determinants of technical efficiency (Battese and Coelli, 1995). The idea of the frontier production function can be illustrated with a farm using n inputs (X_1, X_2, \dots, X_n) to produce output Y . Efficient transformation of

inputs into output is characterized by the production function $f(x)$, which shows the maximum output obtainable from various input vectors. The stochastic frontier production function assumes the presence of technical inefficiency of production. Hence, the function is defined by,

$$Y_i = f(x_i, \beta) \exp(v_i - u_i) \quad i = 1, 2, \dots, n \quad (1)$$

Where: v is a random error associated with random factors not under the control of farmers. The model is such that the possible production Y_i is bounded above by the stochastic quantity $f(x_i, \beta) \exp(v_i)$, hence the term "stochastic frontier". The random error v_i are assumed to be independently and identically distributed as $N(0, \sigma^2 v)$ random variables independent of the u_i s.

The technical efficiency of an individual farmer is defined in terms of the ratio of the observed output to the corresponding frontier output, given the available technology.

$$\begin{aligned} \text{Technical efficiency (TE)} &= Y_i / Y_i^* \\ &= f(x_i, \beta) \exp(v_i - u_i) / f(x_i, \beta) \exp(v_i) \\ &= \exp(-u_i) \end{aligned} \quad (2)$$

Where Y_i is the observed output and Y_i^* the frontier output. Technically efficient farms are those that operate on the production frontier and the level by which a farm lies below its production frontier is regarded as the measure of technical inefficiency. For this study, the production technology of small scale food crop farmers is assumed to be specified by the Cobb-Douglas frontier production function defined by,

$$\ln Y = \beta_0 + \beta_1 \log X_1 + \beta_2 \log X_2 + \beta_3 \log X_3 + \beta_4 \log X_4 + \beta_5 \log X_5 + V_i - U_i \quad (3)$$

Where

\ln represents the natural logarithm

Y represents the value of production of i -th farmer measured in Naira¹

X1 represents the total area of land in hectares on which crops were grown

X2 represents family labour in man days

X3 stands for the value of implements in Naira

X4 represents the quantity of fertilizer used, in kilograms

X5 stands for value of seed in Naira

β is are coefficients to be estimated

V_{is} is assumed to be independent and identically distributed normal random errors, having zero mean and unknown variance, σ^2v ;

The U_{is} are the technical inefficiency effects, which are assumed to be independent of V_{is} such that U_{is} is the non-negative truncation (at zero) of the normal distribution with mean, μ_i , and variance, σ^2 , where μ_i is defined by,

$$\mu_i = \alpha_0 + \alpha_1 z_{1i} + \alpha_2 z_{2i} + \alpha_3 z_{3i} + \alpha_4 z_{4i} + \alpha_5 z_{5i} \quad (4)$$

Where: z_1, z_2, z_3, z_4, z_5 are age, level of education, farming experience, farm size and family size of farm operator, respectively. These variables are assumed to influence the technical efficiency of the farmers and are the unknown scalar parameters to be estimated. The variables age, level of education, farming experience, farm and family sizes are included in the model as determinants of technical efficiency to indicate the possible effects of farmers' characteristics on technical efficiency in order to be able to come out with recommendations on how government policy formulation could be used to influence these variables so as to enhance the technical efficiency of the farmers.

Discussion

The Results of Descriptive Statistics

The result of descriptive statistics of variables used in the stochastic frontier production function is as presented below. The values in the summary statistics vary across the two zones. The farmers involved in the study have relatively small farms. Farm sizes for both zones ranged between 0.493 and 2.20 hectares. Also, both hired and family labour was extensively used by the

respondents, though with wide variations within and between zones. The main reason for the wide variation in the intensity of farm labour use could be attributed to variation in the types of crops grown by respondent farmers. For instance, yam production is known to be traditionally associated with intensive labour use, especially with mould-making, staking and other operations involved in yam farming.

Table 1: Summary of Descriptive statistics of variables in the stochastic frontier model for the small scale farmers.

Variables	Mean	Standard Deviation	Minimum
	Maximum	Coeff of variation	
Value of output (Naira)	28,303	39,199	1395
Farm size (Hectares)	74250	1.38	
Total Labour (Mandays)	1.72	0.493	0.900
Hired Labour (Mandays)	2.20	0.29	
Value of seed (Naira)	90	28.9	17
Implements (Naira)	201	0.32	
Fertilizers (Kg)	39	50	8
Age (years)	104	1.28	
Education (years)	500	205.7	127
Farming Experience	871	0.41	
Family size	400.2	534.76	140
	1,536	1.33	
	52	38	21
	300	0.73	
	38	5.9	-21
	70	0.16	
	4	6.2	0
	12	1.55	
	19	4.9	4
	28.5	0.26	
	6	3.7	1
	10	0.62	

Source: Field Survey, 2012.

The Results of Maximum Likelihood Estimates

The inferences about the stochastic frontier model on the maximum likelihood estimates are represented by the elasticity estimates. The variance parameters of the model is obtained in terms of:

$$\alpha_2s = \alpha u^2 + \alpha v^2 \text{ and } \alpha_2 / (\alpha v^2 + \alpha u^2)$$

The estimate for the parameter in the stochastic frontier model (87%) is quite large. The value indicates the relative magnitude of the variance with the inefficiency effects. This implies that technical inefficiencies are highly significant in the analysis of the data. The production elasticity measures the proportional change in output, resulting from a proportional change in the i-th input level, with all other input levels held constant. Presented in Table 2 are elasticity estimates and returns-to scale value.

Table 2: Elasticity and Returns-to-Scale for Small Scale Farmers in Kogi State

Land	0.23
Labour	0.34
Implements	0.27
Agrochemicals	0.18
Seeds	0.24
Returns-to-scale	1.26

Source: Field Survey, 2012

The elasticity of the mean values of output with respect to the inputs are estimated at the values of the means of the resources. The elasticity of the mean value of farm output with respect to land, labour, implements, agrochemicals and seeds are 0.23, 0.34, 0.27, 0.18 and 0.24, respectively. Given the specification of the Cobb-Douglas frontier models, the results show that the elasticity of the mean value of farm output is estimated to be an increasing function of land, an increasing function of labour, and an increasing function of implements. Also, the mean value of farm output is estimated to be an increasing function of agrochemicals, as well as that of seeds. The returns-to-scale value, 1.26, indicates an increasing returns-

to- scale. The returns-to-scale parameter indicates what happens when all production resources are varied in the long run by the same proportion. The implication of increasing-returns-scale in this study means increasing productivity per unit of input. The farmers are not using their resources efficiently. They can still increase their level of output at the current level of resources.

Technical Efficiency Estimates

Given the specification of the Cobb-Douglas stochastic frontier model in equation (1), the predicted technical efficiency differs widely among the sample farmers, with minimum and maximum values of 0.19 and 0.93, respectively and a mean technical efficiency value of 0.64. Table 4 presents the frequency distribution of the technical efficiency of the sampled farmers.

Table 4: The Frequency Distribution of Technical Efficiency Estimates.

Efficiency Range	Frequency	% of Total
0.1 – 0.29	15	7.5
0.30 – 0.49	43	21.5
0.50 – 0.69	106	53.0
0.70 – 0.89	31	15.5
0.90 – 1.00	5	2.5
Total	200	100

Source: Field Survey, 2012

The distribution of the technical efficiency in table 3 clearly shows that technical efficiency skewed heavily in the 0.50 and 0.69 range, representing 53% of the sampled farmers. The wide variation in technical efficiency estimates is an indication that most of the farmers are still using their resources inefficiently in the production process and there still exist opportunities for improving on their current level of technical efficiency. Given the results of the inefficiency model in the Cobb-Douglas frontier model, the age of operator, the level of education and the farming experience of operators are individually significant determinants of technical inefficiency at 5% level. The implication here is that these variables significantly affect the level of the technical efficiency of the respondent farmers. However, family size and farm size did not significantly influence technical inefficiency. While the level of education, farm size and farming experience have negative coefficients; the age of the operator and family sizes have

positive coefficients, respectively. The negative coefficients of level of education, farm size and farming experience imply that an increase in any of or in all of these variables would lead to a *decline* in the level of technical inefficiency. An increase in the value of variables with positive coefficients (age of operator and family size) implies that an increase in the value of these variables would lead to an *increase* in the level of technical inefficiency. In order to determine the magnitude of change in the level of technical efficiency, that could result as a result of change in government policies that influence the determinants of technical inefficiency, a simulation analysis was performed on the identified variables, which could be influenced by government policy.

Analysis of Policy Variables that Affect Technical Inefficiency

The simulation result is shown in Table 4 assuming a change in policy that influences the determinants of technical inefficiency. The simulation is done with an increase in the values of the variables by 5%, 10% and 20% and the observed change in the level of technical efficiency is as presented below.

Table 4: Simulation Results of Variation in Policy Variables on Mean Technical Efficiency.

Variable	Mean T.E = 0.63		
	+5	+10	+20
Age of Operator	0.65	0.64	0.63
Level of Education	0.67	0.69	0.72
Farming Experience	0.67	0.68	0.71
Family size	0.65	0.65	0.64
Farm size	0.67	0.68	0.69

Source: Computed from Field Survey 2012

The results of the simulation of policy variables show that the mean technical inefficiency would *decline* with rising levels of education, farming experience and farm size. An increase in the level of education from 5% through 20% *raised* the mean technical efficiency from the current level of 67% to 72%, while an increase in the level of farming experience from 5% through 20% led to *increase* in the mean technical efficiency from the current level of 67% to 71%. On the other hand, an

increase in farm size from 5% through 20% only led to a *marginal increase* in the mean technical efficiency. An increase in the age and the family size of operator from 5% through 20% led to a significant *decline* in the mean technical efficiency from 65% to 63% and from 65% to 64%, respectively. The implication of the foregoing analyses is that education is one of the policy variables which can be used by policy makers to improve the current level of technical efficiency of farmers in Nigeria. Hence, any agricultural policy in the country that would attract people with high level of education into farming and/or encourage illiterate farmers to undergo education/training would definitely lead to an increase in the levels of technical efficiency of the farmers. Also, the analyses imply that any agricultural policy in the country that would encourage experienced farmers to remain in the farming business (thereby gaining more experience) would also lead to an increase in the level of technical efficiency of the farmers. It is also important to state that any agricultural policy that would attract young people into the farming business would lead to an increase in the level of technical efficiency, given that young and educated people are more receptive to agricultural innovation than old and illiterate farmers. The current government policy, which encourages a maximum of four children per woman will in the long run lead to a decline in family size, especially among the farming families such a decline is expected to result in an increase in the level of technical efficiency (Table 4), given that the farmers have small farm sizes and most family members are underemployed on the farm.

Conclusion

In conclusion, the education level of farmers and farming experience are important policy variables and determinants of efficiency, which can be incorporated into the agricultural policy in Kogi State in order to raise the current level of technical efficiency and hence the level of productivity in the Nigerian agricultural sector.

References

- Abdulkadri, A. O. Ajibefun, I. A. and Ogundare, F. N (1999), "Technical efficiency of Small Scale Oil Palm Processors in Nigeria", *American Journal of Agricultural Economics*, Vol. 81 (5): 1328.
- Adesina, A. A. and K. K. Djato (1997), "Relative Efficiency of Women as Farm Managers: Profit Function Analysis in Cote d'Ivoire", *Agricultural Economics*, Vol. 16, :47-53.

- Aigner, D, C. A. K. Lovell and P. Schmidt (1977), "Formulation and Estimation of Stochastic Frontier Production Models", *Journal of Econometrics*, Vol. 6,: 21 - 37.
- Ajibefun, I. and O. A Abdulkadri (1999), "An Investigation of Technical Inefficiency of Production of Farmers Under the National Directorate of Employment in Ondo State, Nigeria", *Applied Economics Letter*, Vol. 6: 111- 114, Routledge, London.
- Ajibefun, I. A., Battese, G.E. and Daramola, A.G. (1996), "Investigation of Factors Influencing the Technical Efficiencies of Smallholder Croppers in Nigeria", CEPA Working Papers No. 10/96 Department of Econometrics, University of New England, Armidale, Australia, Pp 19.
- Akpakpan, E. B., 1999. *The Economy: Towards a New Type of Economy*, p.25. Belpot, Nigeria
- Bagi, F. S. and C. J. Huang (1983), "Estimating Production Technical Efficiency for Individual Farms in Tennessee" *Canadian Journal of Agricultural Economics*, 31, 249-256.
- Battese, G. E and G. S. Corra (1977), "Estimation of a Production Frontier Model: With Application to the Pastoral Zone of Eastern Australia", *Australian Journal of Agricultural Economics*, Vol. 21: 169-179.
- Brova-Ureta, B. and A. E. Pinheiro (1993), "Efficiency Analysis of Developing Country Agriculture: A Review of the Frontier Function Literature", *Agricultural Resource Economics Review*, Vol. 22: 88-101.
- Battese, G. E and T. J. Coelli (1995), "A Model For Technical Inefficiency Effect in a Stochastic Frontier Production Function for Panel Data", *Empirical Economics*, Vol. 20: 325-332.
- CBN (2010) Central Bank of Nigeria Annual Reports and Statement of Accounts 220-234
- Edeth, U and Oluwatoyin, F (2006), "Resource use efficiency and productivity among farmers in Nigeria", *Journal of Agriculture and Social Sciences* Vol.2 : 264 - 268
- Kumbhakar, S .C, B. Biswas, and D. V. Bailey, (1989), "A study of Economic Efficiency of Utah Dairy Farmers: A System Approach", *Review of Economics and Statistics*, Vol.7:595 -604